

Hyetograph Simulation of High-Intense Rainfall Events

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Abstract. The simulation of storm hyetographs is an important issue for various applications in hydrology and soil sciences. In this study we present a modification of the simple scaling model of Koutsoyiannis and Foufoula-Georgiou (1993), which assume a piecewise linear relationship between the logarithmic duration and logarithmic expected depth of an event. The expectation of the rainfall depth is estimated by local regression smoothing.

1. Introduction

Hyetographs (rainfall-intensity curves) picture the time distribution of point rainfall and are important for hydrologic and hydraulic design issues. In the study of soil erosion, specifically on movement of soil particles by raindrop impact, kinetic energy is a commonly suggested indicator of the raindrop's ability to detach soil particles from the soil mass. The kinetic energy of raindrops can be estimated from rainfall intensity.

Since only rainfall with sufficient intensity is effective, we focus on high-intense rainfall events. From historical rainfall records, we

- i. separate events from rainfall time series (6-hours-criterion)
- ii. use only events with total rain depth $\geq 15\text{mm}$
- iii. use only events with maximum intensity $\geq 10\text{mm/hr}$

2. Summary of the Scaling Model of Storm Hydrograph

The scaling model of storm hydrograph (Koutsoyiannis and Foufoula-Georgiou, 1993) is a stochastic model that describe the time distribution of rainfall intensity and incremental rainfall within a storm event, taking advantage of scaling properties revealed in rainfall data.

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The main hypothesis of the model is that the process of instantaneous rainfall intensities within storms is a self-similar (simple scaling) process with a scaling exponent H . If D denotes the duration of the storm and $\xi(t, D)$ the rainfall intensity at time $t, 0 \leq t \leq D$, then

$$\{\xi(t, D)\} =^d \{\lambda^{-H} \xi(\lambda t, \lambda D)\} \quad (1)$$

where the symbol $=^d$ denotes equality in distribution. A secondary hypothesis is the weak stationarity within the same storm event or within storm events of the same duration.

On the basis of these hypotheses, the moments of the rainfall intensity at any time t within a storm of any duration D can be deduced. The expectation value is

$$E[\xi(t, D)] = c_1 D^H \quad (2)$$

and the product moment

$$E[\xi(t, D)\xi(t + \tau, D)] = \phi(\tau / D) D^H \quad (3)$$

where τ is a time lag, $c_1 = E[\xi(t / D, 1)]$ and $\phi(\cdot)$ is a function to be specified. In this paper we specify this function as .

$$\phi(y) = \kappa y^{-\beta} \quad (4)$$

where β and κ are parameters ($0 < \beta < 1, \kappa > 0$).

The total storm depth over duration D ,

$$Z = \int_0^D \xi(s, D) ds \quad (5)$$

has the mean

$$E[Z] = c_1 D^{H+1} \quad (6)$$

and the variance

$$Var[Z] = c_2 D^{2(H+1)} \quad (7)$$

where c_2 is the variance of the depth of a storm with duration $D = 1$.

Moving from continuous time to discrete time we use the incremental storm depth at the interval $[(i-1)\Delta, i\Delta]$ of length Δ :

$$X_i = \int_{(i-1)\Delta}^{i\Delta} \xi(t, D) dt, \quad i=1, 2, \dots, k \quad (8)$$

where $k = D / \Delta$ is assumed to be an integer. The expectation of X_i is

$$E[X_i] = c_1 \delta D^{H+1} \quad (9)$$

its variance is

$$Var[X_i] = [(c_2 + c_1^2) \delta^{-\beta} - c_1^2] \delta^2 D^{2(H+1)} \quad (10)$$

where $\delta = \Delta / D$. The covariance between two incremental depths within the same storm is given by

$$\text{Cov}[X_i, X_j] = [(c_2 + c_1^2)\delta^{-\beta} f(|j-i|, \alpha, \beta) - c_1^2]\delta^2 D^{2(H+1)} \quad (11)$$

where $f(m, \beta) = (1/2)[(m-1)^{2-\beta} + (m+1)^{2-\beta}] - m^{2-\beta}$. The derivations of the above equations are given by Koutsoyiannis and Foufoula-Georgiou (1993).

The scaling model, in its simplest form, has four parameters: the scaling exponent H , the expectation parameter c_1 , the variance parameter c_2 and the covariance decay parameter β . Taking the logarithms in (6) we get

$$\log E[Z] = \log c_1 + (H+1)\log D. \quad (12)$$

By local regression smoothing we can estimate $E[Z]$ in dependency of D . Thus H and c_1 are directly estimated by least squares from (12), and c_2 from (7). The parameter β is estimated directly from the identity

$$\beta = \frac{\log(E[X_i X_{i+1}] / E[X_i^2] + 1)}{\log 2} \quad (13)$$

which was derived from (9) and (11) in Koutsoyiannis and Pachakis (1996).

3. The Modified Scaling model

In Koutsoyiannis and Foufoula-Georgiou (1993) a global validity of the linear relationship of $\log E[Z]$ and $\log D$ is taken for granted. But scatterplots like Fig. 1 indicate a more or less significant deviation of the linear relationship.

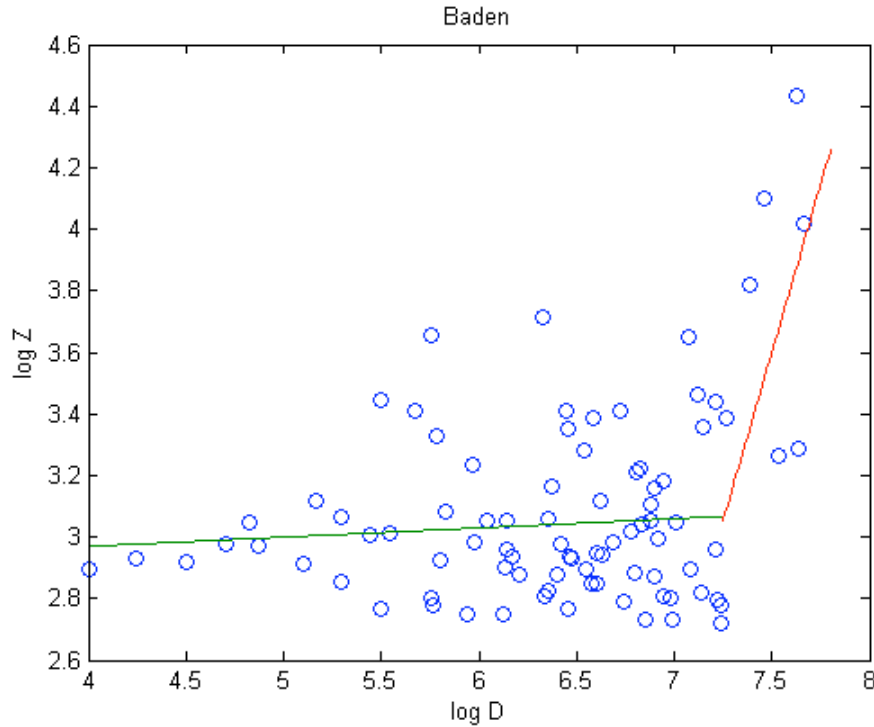


Figure 1. Scatterplot and piecewise-linear model

In surprisingly many cases a single breakpoint is present and the scaling property is (approximately) valid, above and below this breakpoint D^* with different scaling exponents. Typically D^* is located in the range of 19 to 22 hours. Setting $y = \log E[Z]$, $t = \log D$ and $t^* = \log D^*$ we get from (12)

$$\begin{aligned} y_1 &= \beta_{11} + \beta_{12}t, t < t^* \\ y_2 &= \beta_{21} + \beta_{22}t, t > t^* \end{aligned} \quad (14)$$

A piecewise linear regression line has to satisfy the constraint $y_1(t^*) = y_2(t^*)$. The parameters $\beta_{11}, \beta_{12}, \beta_{21}$ and β_{22} in (14) are estimated by least squares method subject to the constraint $y_1(t^*) = y_2(t^*)$. The procedures for least square estimation under constraints can be found in most textbooks on regression analysis, e.g. Draper and Smith (1998).

The modified scaling model has seven parameters: the expectation parameters c_1^{LOW} and c_1^{HIGH} , the variance parameter c_2^{LOW} and c_2^{HIGH} , the scaling exponents H^{LOW} and H^{HIGH} and the covariance decay parameter β . Since the r.h.s. of (13) does not depend on D is β a global parameter.

4. Case Studies

As test cases for the modified scaling model, data sets from three rain gauges were used: a rain gauge at St. Pölten, Lower Austria, one at Steyr, Upper Austria, and a rain gauge at Baden, Lower Austria. The available time resolution is 5 min and the depth resolution is 0.01 mm.

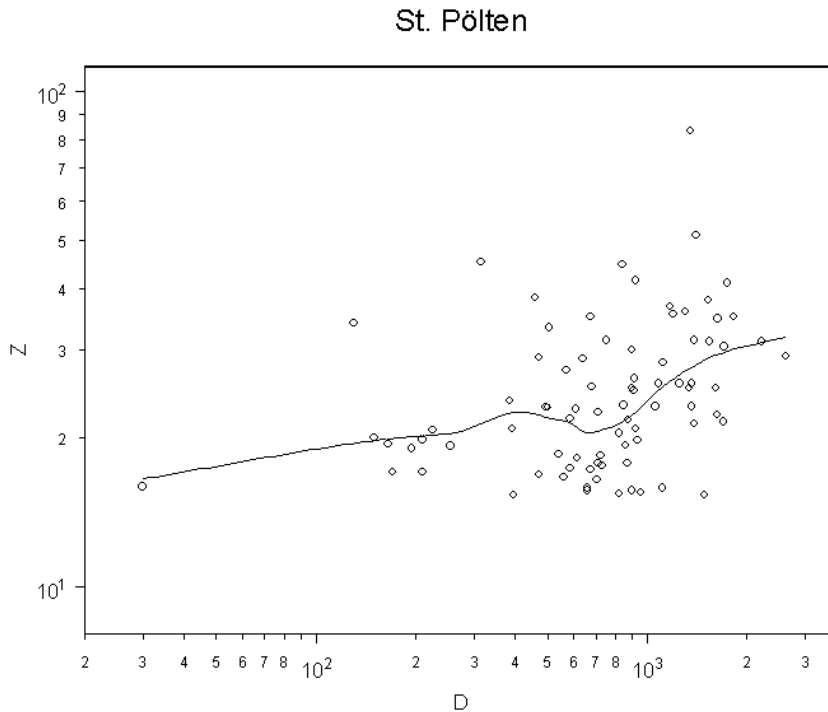


Figure 2. Scatterplot and local regression curve

Using the criterion stated in the introduction, a data set of 164 events were extracted. Fig. 2 shows a scatterplot where $\log Z$ against $\log D$ was plotted. The solid line is a regression curve obtained local regression smoothing. As a smoothing tool we used the routine *loess.smooth* of S-Plus (2001). The curve shown in Fig. 2 exhibits no distinct breakpoint. So the original scaling model with a global scaling exponent can be used. The local regression function estimates for a given duration time D the value of $E[Z]$. This estimation method seems to us more efficient than the rough grouping approach used in Koutsoyiannis and Foufoula-Georgiou (1993). The estimates of $E[Z]$ were inserted in (12) to perform the parameter estimation procedures described in section 2.

For the data set St. Pölten the following estimates were obtained:

$$\hat{H} = -0.85 \quad \hat{c}_1 = 8.75 \quad \hat{c}_2 = 12.13 \quad \beta = 0.27 \quad \hat{\kappa} = 56.0 \quad (15)$$

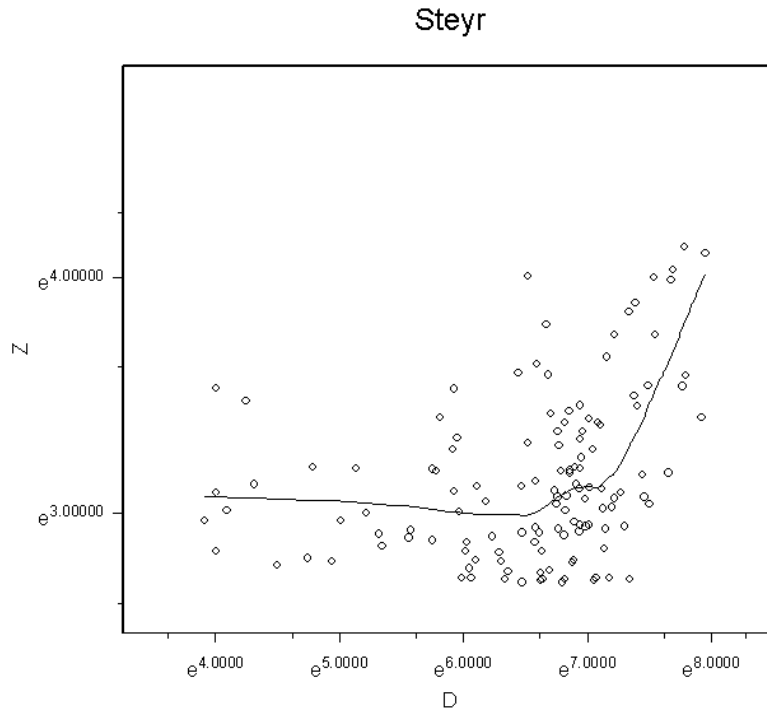


Figure 3. Scatterplot and local regression curve

Fig. 3 depicts the scatterplot of the data set Steyr consisting of 220 storm events. The regression curve exhibits a bend and a breakpoint near 20.8 h. The location of the breakpoint was estimated such that squared deviation of the piece-wise linear model given by (14) from the regression curve is minimal (see Fig. 5). In the range below the breakpoint $\log D^* = 7.13$ the following estimates were obtained:

$$\hat{H} = -0.97 \quad \hat{c}_1 = 18.62 \quad \hat{c}_2 = 34.83 \quad \beta = 0.29 \quad \hat{\kappa} = 231.61, \quad (16)$$

in the range above the breakpoint

$$H = 0.22 \quad \hat{c}_1 = 0.65 \quad \hat{c}_2 = 1.97 * 10^{-3} \quad \beta = 0.29 \quad \hat{\kappa} = 0.26. \quad (17)$$

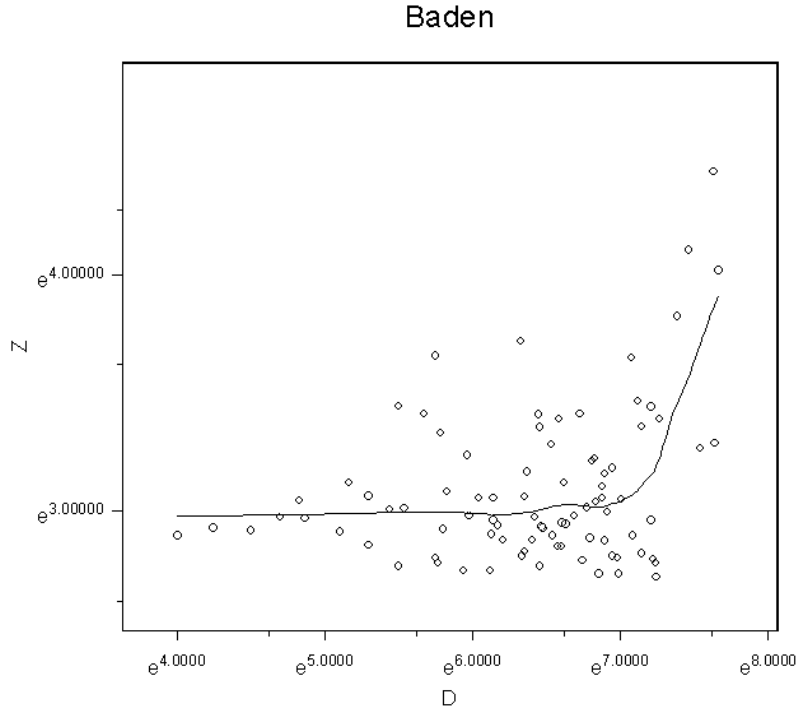


Figure 4. Scatterplot and local regression curve

Fig. 4 depicts the scatterplot of the data set Baden consisting of 220 storm events. A breakpoint was found near 23.2 h. Again a piece-wise linear model was fitted with $\log D^* = 7.24$. Below the breakpoint we obtained the estimates

$$H = -0.98 \quad \hat{c}_1 = 17.28 \quad \hat{c}_2 = 17.36 \quad \beta = 0.29 \quad \hat{\kappa} = 191.80 \quad , \quad (18)$$

above we obtained

$$H = 1.19 \quad \hat{c}_1 = 0.02 \quad \hat{c}_2 = 1.01 * 10^{-4} \quad \beta = 0.29 \quad \hat{\kappa} = 3.04 * 10^{-4} . \quad (19)$$

Fig. 5 and 6 shows the local regression curves and the corresponding piece-wise linear models for the data sets Steyr and Baden, respectively.

5. Simulation of storm hyetographs

The scaling model can be used for simulating storm hyetographs at an incremental basis for any time step Δ . In our simulation studies we used the disaggregation procedure where a given total storm depth Z with a duration D is disaggregated into incremental depths. We have to generate a random vector $\mathbf{X} = (X_1, X_2, \dots, X_k)^T$, where $k = D/\Delta$ is assumed to be an integer. The covariance structure of \mathbf{X} is given by the scaling model, cf. formula (11). Moreover we have to specify the marginal distribution. For rainfall depths the Gamma distribution is the first choice. Let $\mathbf{Z} = (Z_1, Z_2, \dots, Z_k)^T$ be a k -vector of two-parameter Gamma distributed random variables with

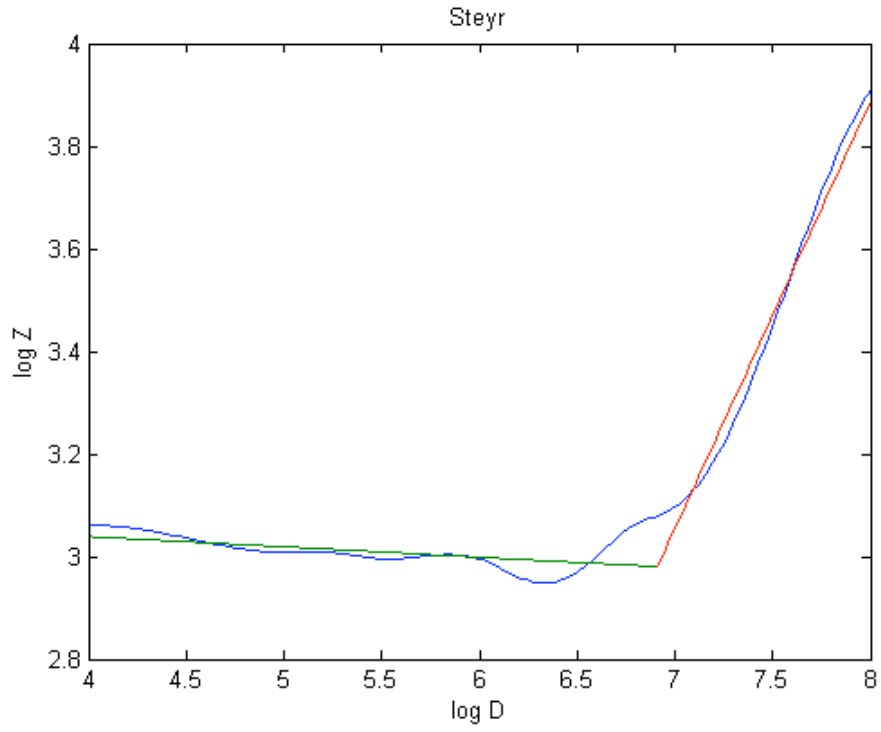


Figure 5. Local regression curve and piecewise-linear model

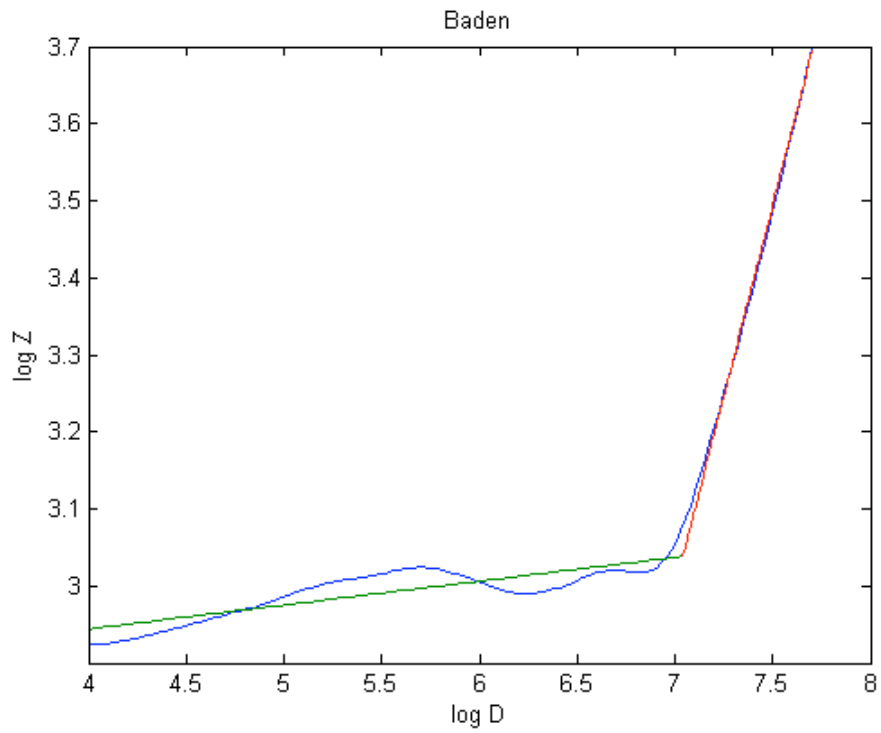


Figure 6. Local regression curve and piecewise-linear model

identity covariance matrix. Sim (1993) proposed a procedure to generate \mathbf{X} by a linear transformation

$$\mathbf{X} = \mathbf{B} \mathbf{Z}. \quad (20)$$

\mathbf{B} is a lower triangular matrix of Beta distributed random variables. Some details of Sim's method are given in the appendix. Different from the sequential process proposed by Koutsoyiannis and Foufoula-Georgiou (1993), Sim's procedure results always non-negative exactly Gamma distributed output.

Making use of the scale invariance the total storm is disaggregated by the following process:

- i. Apply Sim's procedure to obtain an initial sequence $Y_i, i = 1, \dots, k$.
- ii. Determine a normalized sequence

$$Y_i^* = \frac{Y_i}{\sum_{j=1}^k Y_j}. \quad (21)$$

- iii. Calculate the final sequence $X_i = Y_i^* Z$.

A disadvantage of Sim's procedure is that the given covariance matrix Σ must have positive entries, that is all the components of the random vector \mathbf{X} must be positively correlated.

6. Conclusions

A modification of the simple scaling model has been presented, with different scaling exponents for small and large storm durations. The expected storm depth, as a function of storm duration, was estimated by local regression smoothing. A piecewise linear model was fitted by least squares estimation. So the scaling model can be used even in the case of nonlinear relationship of $\log E[Z]$ and $\log D$. The presence of a single breakpoint was observed in many data sets of storm events.

Koutsoyiannis and Mamassis (2001) compared the simple scaling model with the rectangular pulse model of Rodriguez-Iturbe et al. (1987) applied to a single storm hyetograph. They found that the scaling model reproduces the structure of historical hyetograph better than the rectangular pulse model. Recently Cowpertwait et al. (2007) proposed a refined rectangular pulse model with an appropriate fine-scale structure. A comparison of the modified scaling model with a refined rectangular pulse model will be the subject of a further study.

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Appendix

In this appendix we give a short description of Sim's generation process, specialized to our simulation problem. Let $\mathbf{Z} = (Z_1, Z_2, \dots, Z_k)^T$ be a k-dimensional random vector with identity covariance matrix and $\text{Gamma}(r_i, \beta)$ marginal distributions. Let $B_{ji}, 1 \leq i < j \leq k$, be independent $\text{Beta}(a_{ji}, r_i - a_{ji})$ random variables which are independent of Z_i , and $0 < a_{ji} < r_i$. The required gamma random vector $\mathbf{X} = (X_1, X_2, \dots, X_k)^T$ which has given $\text{Gamma}(\alpha, \beta)$ marginal distributions and a given covariance matrix $\Sigma = (\sigma_{ij})$, where $\sigma_{ii} = \alpha / \beta^2$, can be obtained via the transformation (20), with $B_{ii} = 1$. The random variables $B_{ji}Z_i$ have $\text{Gamma}(a_{ji}, \beta)$ marginal distributions. The aim is to determine the coefficients a_{ji} and parameters r_i given α, β and Σ via the recursive formula

$$a_{ji} = \beta^2 \sigma_{ij} - \sum_{p=1}^{i-1} (a_{ip} a_{jp} / r_p) \quad (22) \quad \text{and}$$

$$r_j = \alpha - \sum_{p=1}^{j-1} a_{jp} . \quad (23)$$

Note we have to ensure, that $0 < a_{ji} < r_i, 1 \leq i < j \leq k$. This conditions are required by the construction of (20). It means that this simulation procedure will stop if any of these parameters goes beyond these intervals.

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